

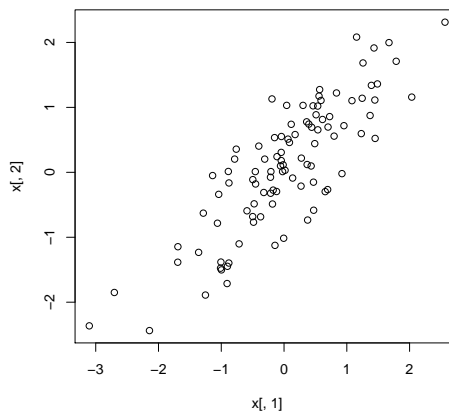
Principal Component Analysis

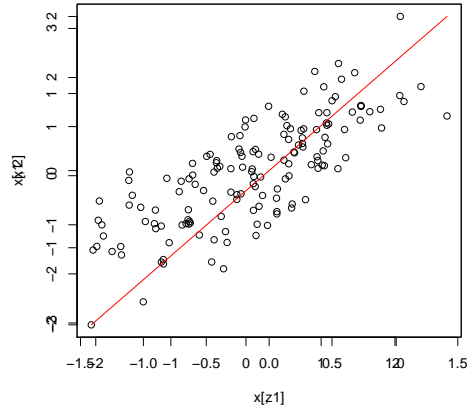
1 Introduction

Principal component analysis is a dimension reduction technique in multivariate analysis. The principal idea of reducing the dimension of $X = (X_1, X_2, \dots, X_p)'$ is achieved through linear combination. Low dimensional linear combinations are easier to interpret and serve as an intermediate step as in a more complex analysis. More precisely one looks for linear combination which create largest spread among the values of X . In other words, one is searching for linear combinations of variables with largest variances.

Consider an the following example

Let $(X_1, X_2)' \sim N_2((0, 0)', \Sigma)$ where $\Sigma = \begin{pmatrix} 1 & 0.8 \\ 0.8 & 1 \end{pmatrix}$. The scatterplot of 100 sample drawn from this distribution is as follows





Let us consider the following transformation $\mathbf{X} \rightarrow \mathbf{Y}$

$$Y_1 = a_{11}X_1 + a_{12}X_2$$

$$Y_2 = a_{21}X_1 + a_{22}X_2$$

or

$$\mathbf{Y} = \mathbf{A}\mathbf{X}$$

The components of A are chosen such that $Var(Y_1)$ is maximized. The vectors $a'_i = (a_{i1}, a_{i2})'$ for $i = 1, 2$ are chosen such that $a_i a'_j = 0$ and $a_i a'_i = 1$, that is the row vectors of coefficient matrix are orthonormal.

Let

$$Y_1 = \frac{1}{\sqrt{2}}X_1 + \frac{1}{\sqrt{2}}X_2$$

$$Y_2 = \frac{1}{\sqrt{2}}X_1 - \frac{1}{\sqrt{2}}X_2$$

Then

$$V(Y_1) = \frac{1}{2} (V(X_1) + V(X_2) + 2Cov(X_1, X_2)) = 1.8$$

and

$$V(Y_2) = \frac{1}{2} (V(X_1) + V(X_2) - 2Cov(X_1, X_2)) = 0.2.$$

So the linear combination Y_1 explains 90% of total variability $V(Y_1) + V(Y_2)$. Y_1 is first principal component. It can be verified that

$$V(X_1) + V(X_2) = V(Y_1) + V(Y_2)$$

General Model

$$\begin{aligned} Y_1 &= a_{11}X_1 + a_{12}X_2 + \dots + a_{1p}X_p = \mathbf{a}'_1 \mathbf{X} \\ Y_2 &= a_{21}X_1 + a_{22}X_2 + \dots + a_{2p}X_p = \mathbf{a}'_2 \mathbf{X} \\ &\vdots \\ Y_p &= a_{p1}X_1 + a_{p2}X_2 + \dots + a_{pp}X_p = \mathbf{a}'_p \mathbf{X} \end{aligned}$$

or

$$\mathbf{Y} = \mathbf{A}\mathbf{X}$$

$\mathbf{Y} = (Y_1, Y_2, \dots, Y_p)'$ are the principal components. Y_j is the j^{th} principal component.

The model is defined once the elements of A are estimated. Firstly $\mathbf{a}'_1, \mathbf{a}'_2 \dots \mathbf{a}'_p$ needs to be orthogonal that is $\mathbf{a}'_j \mathbf{a}_k = 0$ and $\mathbf{a}'_j \mathbf{a}_j = 1$.

Choose \mathbf{a}_1 such that $Var(Y_1) = Var(\mathbf{a}'_1 \mathbf{X})$ is maximized with respect to \mathbf{a}_1 . \mathbf{a}_2 such that $Var(Y_2) = Var(\mathbf{a}'_2 \mathbf{X})$ is maximized with respect to \mathbf{a}_2 subject to the orthogonality condition and so on. Here we have $cov(Y_j, Y_k) = 0$ and

$$Var(Y_1) \geq Var(Y_2) \geq \dots \geq Var(Y_p)$$

Also,

$$\sum_{i=1}^p Var(X_i) = \sum_{i=1}^p Var(Y_i).$$

For the first $Y_1, Y_2, \dots, Y_q, q < p$ we compute

$$\frac{\sum_{i=1}^q Var(Y_i)}{\sum_{i=1}^p Var(Y_i)} \times 100$$

If it is reasonably large we can consider these q variables instead of p original ones.

2 Population principal components

Let $\mathbf{X} = (X_1, X_2, \dots, X_p)'$ be a p variate random vector with $E(\mathbf{X}) = \mu$ and known covariance matrix Σ . We shall consider cases which Σ is positive semidefinite matrix. Since we shall be concerned with variances and covariances of X we shall assume that $\mu = 0$. The first principal component the normalized linear combination $Y_1 = \alpha'X$ where $\alpha = (\alpha_1, \alpha_2, \dots, \alpha_p)'$ is such that $\alpha'\alpha = 1$ and

$$Var(\alpha'\mathbf{X}) = \max_l Var(l'\mathbf{X})$$

with $l' \in \mathbb{R}^p$ satisfying $L'L = 1$.

Now

$$V(l'X) = l'\Sigma l.$$

Thus to find first principal component $\alpha'X$ we need to the α such that maximizes $l'\Sigma l$ for all choices of $l \in \mathbb{R}^p$ subject to the restriction $l'l = 1$. Using the Lagrange's multiplier λ we need to find the α that maximizes

$$\phi_1(l) = l'\Sigma l - \lambda(l'l - 1)$$

for all choices of $l \in \mathbb{R}^p$ satisfying $l'l = 1$. Now

$$\begin{aligned} \frac{\partial \phi_1}{\partial l} &= 0 \\ \text{or } 2\Sigma l - 2\lambda l &= 0 \\ \text{or } (\Sigma - \lambda I)l &= 0 \end{aligned}$$

So, α satisfies the equation $(\Sigma - \lambda I)\alpha = 0 \dots (1)$. Since $\alpha \neq 0$ as a consequence of $\alpha'\alpha = 1$, the equation (1) has a solution if $|\Sigma - \lambda I| = 0$.

That is λ is the characteristic root of Σ and α is the corresponding characteristic vector. Since dimension of Σ is $p \times p$ there are p values of λ . Let

$$\lambda_1 \geq \lambda_2 \geq \dots \geq \lambda_p$$

denote the ordered characteristic roots and

$$\alpha_1 = (\alpha_{11}, \dots, \alpha_{1p})' \dots \alpha_p = (\alpha_{p1}, \dots, \alpha_{pp})'$$

denote the characteristic vectors of Σ . Now, Σ may have zero characteristic root or some the roots may have multiplicity greater than unity.

Now, $(\Sigma - \lambda I) \alpha = 0$ gives

$$Var(\alpha'X) = \alpha'\Sigma\alpha = \lambda\alpha'\alpha = \lambda.$$

where λ is the characteristic root Σ corresponding to α . Thus to maximize $Var(\alpha'X)$ we need to choose $\lambda = \lambda_1$, the largest characteristic root of Σ and $\alpha = \alpha_1$ is the characteristic root corresponding to λ_1 .

So the first principal component is the normalized linear function

$$Y_1 = \alpha_1'X = \sum_{i=1}^p \alpha_{1i}X_i$$

where α_1 is the normalized characteristic vector of Σ corresponding to largest characteristic root λ_1 is called first principal component of X .

The second principal component is the normalized linear function $\alpha'X$ having maximum variance among all normalized linear functions $l'X$ that are uncorrelated with Y_1 .

$$\begin{aligned} cov(l'X, Y_1) &= E(l'XY_1) = E(l'X\alpha_1') = E(l'X\alpha_1) \\ &= l'\Sigma\alpha_1 = l'\lambda_1\alpha_1 = \lambda_1 l'\alpha_1 = 0 \end{aligned}$$

This implies l and α_1 are orthogonal. The second principal component will be linear combination $\alpha'X$ that has maximum variance among all normalized linear combination $l'X, l \in \mathbb{R}^p$ which is uncorrelated with Y_1 . Again by using Lagrange's multiplier we will get the second principal component as

$$Y_2 = \alpha_2'X$$

where α_2 is the characteristic vector corresponding to characteristic root λ_2 .

Continuing in this way we will get the r th principal component as

$$Y_r = \alpha_r'X$$

where α_r is the characteristic vector corresponding to characteristic root λ_r .

Now define matrices

$$A = (\alpha_1, \alpha_2, \dots, \alpha_p), \Lambda = \begin{pmatrix} \lambda_1 & 0 & \dots & 0 \\ 0 & \lambda_2 & \dots & 0 \\ \vdots & \vdots & \dots & \vdots \\ 0 & 0 & \dots & \lambda_p \end{pmatrix}$$

where $\lambda_1 \geq \lambda_2 \geq \dots \geq \lambda_p$ are ordered characteristic roots and $\alpha_1, \dots, \alpha_p$ are the corresponding normalized characteristic vectors. Since $AA' = I$ and $\Sigma A = A\Lambda$ we conclude that $A'\Sigma A = \Lambda$.

So there exists an orthogonal transformation

$$Y = A'X$$

such that $D(Z) = \Lambda$ a diagonal matrix with diagonal elements $\lambda_1 \geq \lambda_2 \geq \dots \geq \lambda_p \geq 0$ the ordered roots of $|\Sigma - \lambda I| = 0$.

3 Sample principal component

In practice the covariance matrix is usually unknown. If sample observations on a multivariate random vector is given we have to replace Σ by an estimate of covariance matrix $\hat{\Sigma}$. Now we assume that $\mathbf{X} \sim N_p(\mu, \Sigma)$ where Σ is positive definite matrix.

Let $x^\alpha = (x_{\alpha 1}, x_{\alpha 2}, \dots, x_{\alpha p})'$, $\alpha = 1, 2, \dots, N$, ($N > p$) be a sample of size N from the distribution of X which is univariate normal with mean vector μ and dispersion matrix Σ .

Let

$$\bar{\mathbf{x}} = \frac{1}{N} \sum_{\alpha=1}^N x^\alpha \quad \mathbf{s} = \sum_{\alpha=1}^N (x^\alpha - \bar{x})(x^\alpha - \bar{x})'$$

The maximumlikelihood estimate of Σ is $\frac{\mathbf{s}}{N}$ and that of μ is $\bar{\mathbf{x}}$.

Theorem: The maximum likelihood estimates of ordered characteristic roots $\lambda_1, \lambda_2, \dots, \lambda_p$ of Σ and corresponding characteristic vector $\alpha_1, \alpha_2, \dots, \alpha_p$ are, respectively the ordered characteristic roots r_1, r_2, \dots, r_p and characteristic vector a_1, a_2, \dots, a_p of \mathbf{s}/N .

Proof Omitted

The estimate of total system variance is given by

$$\text{trace} \left(\frac{\mathbf{s}}{N} \right) = \sum_{i=1}^p r_i$$

and is called total sample variance. The importance of i^{th} principal component is measured by

$$\frac{r_i}{\sum_{i=1}^p r_i}.$$

If the estimates of the principal components are obtained by from sample correlation matrix

$$R = (r_{ij}) \quad r_{ij} = \frac{s_{ij}}{(s_{ii}s_{jj})^{1/2}}$$

with $s = (s_{ij})$ then the total sample variance will be $p = \text{trace}(R)$.

If first k principal components explain large amount of total sampel variance, they may be used in place of original vector \mathbf{X} .

Exercise

1. Let

$$\begin{pmatrix} X_1 \\ X_2 \\ X_3 \end{pmatrix} \sim N_3 \left(\begin{pmatrix} 0 \\ 0 \\ 1 \end{pmatrix}, \begin{pmatrix} 1 & 2 & 1 \\ 2 & 5 & 2 \\ 1 & 2 & 2 \end{pmatrix} \right)$$

Obtain the first principal component and obtain the percentage of variability explained by first principal component.

2. Let $(X_1, X_2, X_3)'$ is trivariate random vector with correlation matrix

$$\begin{pmatrix} 1 & 0.8944 & 0.7071 \\ 0.8944 & 1 & 0.6325 \\ 0.7071 & 0.6325 & 1 \end{pmatrix}$$

Find the first principal component and obtain the percentage of variability explained by first principal component.

3. Find principal components of the data “mtcars” in R. (Use “prcomp” command)